22731

1.

A 120 MINUTES

Poisson process B) Brownian motion process A) Both A and B None of these C) D) 2. In an irreducible Markov chain: A) All states are transient B) All states are persistent C) Some states are transient and others are persistent D) Either all states are transient or all states are persistent 3. Consider two independent series of events A and B occurring in accordance with Poisson process with mean λt and μt respectively. Then the number N of occurrences of A between two successive occurrences of B has: Geometric distribution Exponential distribution A) B) C) Uniform distribution D) **Binomial distribution** 4. Arrivals at a telephone booth are considered to be Poisson with an average time 10 minutes between one arrival and the next. The length of phone calls is assumed to be distributed exponentially with a mean of 3 minutes. What is the probability of that a person arriving at the booth will have to wait? 03 D) 0 A) 07 B) 0.03 C) 5. Yule-Furry process is an example for a: Birth and death process Pure death process A) B) Birth immigration process C) D) Pure birth process 6. In time-series analysis, which source of variation can be estimated by the ratio-totrend method? Trend Cyclical variation A) B) Seasonal variation C) Irregular variation D) 7 In the measurement of secular trend, method of moving average: Measure the seasonal variation A) B) Smooth out the time series C) Give the trend in a straight line None of these D) 8. All the index numbers are affected by: Formula error A) B) Sampling error C) Homogeneity error D) All the above 9. Simple aggregative type of index number satisfies: A) Time reversal and factor reversal tests Time reversal and circular tests B)

Which of the following process is a process with independent increments

- C) Factor reversal and circular tests
- D) None of the three tests

10.	Whic 1. 2. A) C)	 ch of the following statements are true? A set may have no limit point, unique limit point or any finite or infinite number of limit points limit point of a set may or may not be a member of the set 1 only B) 2 only Both 1 and 2 D) Neither 1 nor 2 							
	,	Both 1 and 2			D)				
11.	The l	imit points of t	he set {	1, -1, 1	$\frac{1}{2}, -1\frac{1}{2}$	$,1\frac{1}{3},-1$	$1\frac{1}{3}, \dots$ is/are	:	
	A) C)	Does not exit (-1, 1)	ist		B) D)	Only Only	1 and -1 0		
12.	If <i>f</i> (:	$x) = \begin{cases} 1, & \text{if } x \text{ is} \\ 0, & \text{if } x \text{ if} \end{cases}$	s irrati is ratio	onal nal ^{, th}	nen the v	alue of	$f(f \circ f)(\sqrt{3})$ is	3:	
	A)	0	B)	1		C)	$\sqrt{3}$	D)	3
13.	Whic 1. 2. 3.	th of the follow The set of al The set of a Every count	l rationa 11 irratio	al numb mal nun	ers nbers	ero?			
	A) C)	1 and 2 only 2 and 3 only			B) D)	1 and 3 only	-		
14.		f $f(x) = \begin{cases} x + x^2, when x is rational \\ x^2 + x^3, when x is irrational \end{cases}$ then the value of upper Reimann integral n (0, 2) is:							
	A)	0	B)	53 12		C)	12	D)	83 12
15.	If ν is a signed measure and $E_1 \subseteq E_2 \subseteq \cdots$, then $\nu(\bigcup_{i=1}^{\infty} E_i) =$								
	A)	$\lim \nu(E_i)$	B)	$\nu(\bigcap_{i=1}^{\infty}$	$\sum_{i=1}^{5} E_i$	C)	$\sum_{i=1}^{\infty} \nu(E_i)$	D)	None of these
16.	Gene A) C)	· –							
17.	Let $X = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}$ be a random vector with mean vector $\mu = \begin{bmatrix} 5 \\ 3 \end{bmatrix}$ and variance covariance matrix $\Sigma = \begin{bmatrix} 2 & 1 \\ 1 & 2 \end{bmatrix}$. Then the covariance matrix of $Z_1 = X_1 - X_2$ and $Z_2 = X_1 + X_2$ is:								
	A)	$\begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}$	B)	$[^{2}_{1}$	1 2]	C)	$\begin{bmatrix} 0 & 2 \\ 2 & 0 \end{bmatrix}$	D)	$\begin{bmatrix} 4 & 1 \\ 1 & 4 \end{bmatrix}$

18. Let
$$X = \begin{bmatrix} X_1 \\ - \\ X_2 \end{bmatrix}$$
 be distributed as $N_p(\mu, \Sigma)$ with $\mu = \begin{bmatrix} \mu_1 \\ - \\ \mu_2 \end{bmatrix}$, $\Sigma = \begin{bmatrix} \Sigma_{11} & | & \Sigma_{12} \\ - & | & - \\ \Sigma_{21} & | & \Sigma_{22} \end{bmatrix}$ and

 $|\Sigma_{22}| > 0$. Then the conditional distribution of X₁, given that X₂ = x₂ is:

- A) Normal having mean μ and covariance Σ
- B) Normal having mean $\mu_1 + \Sigma_{12}\Sigma_{22}^{-1}(x_2 \mu_2)$ and covariance $\Sigma_{11} \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}$
- C) Normal having mean $\mu_1 \Sigma_{12}\Sigma_{22}^{-1}(x_2 \mu_2)$ and covariance $\Sigma_{11} \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}$
- D) Normal having mean $\mu_1 + \Sigma_{12}\Sigma_{22}^{-1}(x_2 \mu_2)$ and covariance $\Sigma_{22} \Sigma_{12}\Sigma_{11}^{-1}\Sigma_{21}$
- 19. Let $X_1, X_2, ..., X_n$ be a random sample of size n from $N_p(0, \Sigma)$. Then the distribution of $\sum_{j=1}^n X_j X_j$ is:
 - A) Chi square distribution with n degrees of freedom
 - B) Wishart distribution with n-1 degrees of freedom
 - C) Chi square distribution with n-1 degrees of freedom
 - D) Wishart distribution with n degrees of freedom
- 20. If $\frac{1}{2}$, $\frac{1}{4}$, $\frac{3}{4}$ respectively denote the product moment correlation coefficients between X₁ and X₂, between X₂ and X₃, and between X₁ and X₃, then the multiple correlation of X₁ on X₂ and X₃ is

A)
$$\frac{2}{3}$$
 B) $\sqrt{\frac{2}{3}}$ C) 0 D) 1

- 21. If A and B are any two subspaces of a vector space V over a field F, then which of the following statements are true?
 - 1. A + B is a subspace
 - 2. $A \cap B$ is a subspace
 - 3. $A \cup B$ is a subspace
 - A) 1 and 2 only B) 1 and 3 only
 - C) 2 and 3 only D) 3 only
- 22. The eigen values of the matrix $D = \begin{bmatrix} 0 & 1 & 1 \\ 1 & 0 & 1 \\ 1 & 1 & 0 \end{bmatrix}$ are:
 - A) 0 B) -1, 2 C) 0,1 D)

23. If
$$\rho(A) = rank \ of \ a \ matrix \ A$$
, then which of the following is/are true?
1. $\rho(A) = \rho(A^T)$ 2. $\rho(A) = \rho(AA^T)$ 3. $\rho(A) = \rho(A^TA)$

Does not exist

A)	1 only	B)	1 and 2 only
C)	1 and 3 only	D)	1, 2 and 3

24.	$\begin{vmatrix} 1 + \\ x \\ \vdots \\ x \end{vmatrix}$	$\begin{array}{ccc} x_1 & x_2 \\ x_1 & 1+x_2 \\ & \vdots \\ x_1 & x_2 \end{array}$	$\begin{array}{ccc} \cdots & x\\ \cdots & x\\ \ddots \\ \cdots & 1 + \end{array}$	$\begin{bmatrix} n \\ n \\ \vdots \\ x_n \end{bmatrix} =$					
	A) C)	$ 0 1+x_1+x_2+ $	$\cdots + x_n$		B) D)	1 None	of the above		
25.		uadratic form 2 Positive defin Negative def						ite	
26.	is/are 1.	s the generalize true? G ^T is a gener AG is idempo	alized ir			A, then Y	which of the f	following	statements
		1 only Both 1 and 2			B) D)	2 only Neith	y er 1 nor 2		
27.	Let A the ev A)	ball be drawn f ={0, 1}, B= {(ents A,B, C ar Only pairwis Disjoint	0, 2}, C= e: e indepe	$=\{0, 3\}.$	If four B)	outcor Indep	nes are assum		
28.	Let A and B be mutually exclusive events in the sample space of a random experiment. Suppose the experiment is repeated until either event A or B occurs. Then the probability that the event A occurs before the event B is:								
	A)	$\frac{P(A)}{P(B)}$	B)	$\frac{P(A)}{P(A)+P}$	(B)	C)	$\frac{P(B)}{P(A)+P(B)}$	D)	P(A)P(B)
29.		dice are rolled. a six?	If the tw	vo faces	are dif	ferent,	what is the pr	obability	that at least
30.	infact health	11/36 oratory test re , present. How y persons test bility that a per	ever, the ed. If 1	e test als 0% of	o yield the pop	s a fals pulation	e positive res	ult for 1 j s disease	percent of the
	A)	1/2	B)	1/3		C)	9/11	D)	10/11
31.		joint pdf of $(X_{2})^{1/2} X_{2} = 1 $ is:	X ₁ , X ₂) is	$f(x_1, x_2)$	$(z_2) = \frac{1}{2}$	$x_2 e^{-x_1}$	$x_2, 0 < x_1 < \infty$	∞, 0 < x ₂	$_{2} < 2$, then
	A)	1	B)	0		C)	2	D)	Does not exists

32. Consider a system of n identical components operating independently. Suppose that the length of life of components has common density

$$f(x) = \begin{cases} \frac{1}{\lambda} e^{-x/\lambda}, & x > 0, \ \lambda > 0\\ 0, & otherwise \end{cases}$$

If the components operate in series, then the mean life of the system is:

- A) λ B) $n\lambda$ C) $\frac{\lambda}{n}$ D) $\frac{\lambda^2}{n}$
- 33. Let P(s) be the probability generating function of a nonnegative integer valued random variable X. Then which among the following statements are true:
 - 1 $\frac{d^{k}P(s)}{ds^{k}}\Big|_{s=0} = k! P(X = k)$ 2. $\frac{d^{k}P(s)}{ds^{k}}\Big|_{s=1} = E[X^{k}] \text{, when } E[X^{k}] \text{ exists.}$
 - 3. P(s) does not determine the distribution of X uniquely.
 - 4. If $X_1, X_2, ..., X_n$ be independent random variables, then the probability generating function of sum of X_i 's is the product of probability generating functions of X_i 's.
 - A) 1 and 2 only B) 2 and 3 only C) 3 and 4 only D) 1 and 4 only

34. Let X_1, X_2, \dots, X_n be a random sample taken from $N(\mu, \sigma^2)$. If $\overline{X} = \frac{1}{n} \sum_{i=1}^n X_i$, then the distribution of $X_1 - \overline{X}$ is:

- A) $N\left(0, \frac{n+1}{n}\sigma^2\right)$ B) $N\left(0, \frac{n-1}{n}\sigma^2\right)$ C) $N\left(0, \frac{\sigma^2}{n}\right)$ D) $N\left(0, \frac{2\sigma^2}{n}\right)$
- 35. If *X* follows *t* distribution with 1 degree of freedom then:
 - A) $E(X) = 0 \text{ and } V(X) = \frac{1}{2}$
 - B) E(X) = 0 and V(X) = 1
 - C) E(X) = 0 and V(X) does not exist
 - D) E(X) and V(X) do not exists.
- 36. Let (X, Y) be a bivariate normal (BN) random variable with parameters $\mu_1, \mu_2, \sigma_1^2, \sigma_2^2$ and ρ , and let U = aX + b, $a \neq 0$, V = cY + d, $c \neq 0$. Then the distribution of (U, V) is:
 - A) $BN(a\mu_1, c\mu_2 a^2\sigma_1^2, c^2\sigma_2^2, \rho)$
 - B) $BN\left(a\mu_1 + b, \ c\mu_2 + d, \ a^2\sigma_1^2, \ c^2\sigma_2^2, \ \frac{\rho}{|ac|}\right)$
 - C) $BN(a\mu_1 + b, c\mu_2 + d, a\sigma_1^2, c\sigma_2^2, |ac|\rho)$
 - D) $BN(a\mu_1 + b, c\mu_2 + d, a^2\sigma_1^2, c^2\sigma_2^2, \rho)$

- 37. In a multiple choice oral examination, the grade is based on the number of questions asked until he gets one correct answer. Suppose that a student guesses at each answer and there are 4 choices for each answer. If the trials on assumed to be independent, then the average number of questions required for the first correct answer is: 5 A) 4 B) C) 8 D) 12
- 38. If (X, Y) has trinomial distribution with parameters (n, p_1, p_2) , then the conditional distribution of Y|X = x is:
 - Binomial Hypergeometric A) B) C) Geometric D) Poisson
- 39. An urn contains N marbles numbered 1 through N. Suppose n marble are drawn with replacement. Let M_n be the largest number drawn. Then:

A)
$$P(M_n = k) = \left(\frac{k}{N}\right)^n$$

B)
$$P(M_n = k) = n \left(\frac{k}{N}\right)^{n-1}$$

C)
$$P(M_n = k) = \left(\frac{k}{N}\right)^n - \left(\frac{k-1}{N}\right)^n$$

D)
$$P(M_n = k) = \frac{\binom{k-1}{n-1}}{\binom{n}{N}}, \ k = n, \ n+1, \dots, N$$

40. Suppose the survival times of patients who have advanced cancer of the bladder can be modelled by exponential distribution with mean λ . Then the time by which 25% of the patients will die is:

A)
$$\lambda \log\left(\frac{1}{2}\right)$$
 B) $\lambda \log\left(\frac{4}{3}\right)$ C) $\lambda \log\left(\frac{1}{4}\right)$ D) $\lambda \log\left(\frac{3}{4}\right)$

41. Consider the linear model

$$\underline{\mathbf{y}}_{n\times 1} = A_{n\times p}\beta_{p\times 1} + e_{n\times 1},$$

where $e_{n\times 1} \sim N(0, \sigma^2 I)$. Assume that the rank of A is p. If $l_1'\beta$ and $l_2'\beta$ are the best linear estimates of the estimable functions $l_1'\beta$ and $l_2'\beta$, then the covariance

between $l_1'\beta$ and $l_2'\beta$ is given by:

A)
$$\operatorname{cov}\left(l_{1}'\beta, \ l_{2}'\beta\right) = \sigma^{2}(l_{2}'(A'A)^{-1}l_{1})$$

B) $\operatorname{cov}\left(l_{1}'\beta, \ l_{2}'\beta\right) = \sigma^{2}(l_{1}'(A'A)^{-1}l_{2})$

C)
$$\operatorname{cov}\left(l_{1}\overset{\circ}{\beta}, l_{2}\overset{\circ}{\beta}\right) = \sigma^{2}(l_{1}\overset{\circ}{(A'A)^{-1}}l_{1})$$

D) $\operatorname{cov}\left(l_{1}\overset{\circ}{\beta}, l_{2}\overset{\circ}{\beta}\right) = \sigma^{2}(l_{2}\overset{\circ}{(A'A)^{-1}}l_{2})$

- In a 2^2 design having factors A and B, replicated 4 times, the total of the 42. observations from all replications corresponding to the treatment combinations (1), a, b and ab are respectively -10, -4, -10, and 24. Then the value of the sum of squares due to the interaction effect AB is:
 - 49 C) 100 229.5 A) 28.5 B) D)

- 43. For a symmetric BIBD with parameters (v, b, r, k, λ) , the number of treatments common between any two blocks is:
 - C) $\lambda(r-1)$ b/rA) λ B) $r - \lambda$ D)
- In $p \times p$ Graeco-Latin square design, the degrees of freedom of the error sum of 44. square is equal to:
 - (p-3)(p-2) B) $(p-1)^2$ (p-2)(p-1) D) (p-3)(p-1)A)
 - C)
- 45. What would happen if multiple t-test is performed instead of an ANOVA to compare 10 groups?
 - A) No change in results, except that making multiple comparisons with a t-test requires more computation than doing a single ANOVA
 - B) Making multiple comparisons with a t-test increases the probability making a type I error
 - There is no difference between using ANOVA and using t-test C)
 - D) None of the above
- What effect does increasing the sample size have upon the sampling error? 46.
 - A) It reduces the sampling error
 - It increases the sampling error B)
 - It has no effect on the sampling error C)
 - D) None of the above
- The number of possible samples of size n out of N population size in SRSWR is equal to: A) N^n B) NC_n C) $\frac{N-n}{N}$ D) $\frac{N}{n}$ 47.
- Suppose that a simple random sample of n units is taken from a population of size N. 48. If V_{srswor} and V_{srswr} respectively denote the variance of the sample mean in SRSWOR and SRSWR, then:
 - $V_{srswor} = \frac{(N-n)}{N} V_{srswr} \qquad B) \qquad V_{srswor} = \frac{(N-n)}{(N-1)} V_{srswr}$ $V_{srswor} = \frac{(N-1)}{(N-n)} V_{srswr} \qquad D) \qquad None of the above$ A) C)
- 49. If the sampling frame of the elements are not available, then the sampling technique usually used is:
 - Systematic sampling B) Stratified sampling A)
 - Simple random sampling D) Cluster sampling C)
- 50. In cluster sampling, clusters are formed in such a way that
 - Variation within clusters is minimum while variation between clusters is A) maximum
 - Variation within clusters is maximum while variation between clusters is B) minimum
 - Variation within clusters is minimum while variation between clusters is C) minimum
 - D) Variation within clusters is maximum while variation between clusters is maximum

- 51. Which of the following is a procedure for selecting ppswor sample?
 - A) Cumulative total method
- B) Lahiri's method
- C) Midzuno-Sen method D) All the above

52. Which of the following statement(s) is/are true?

- 1. Ratio estimator make use of auxiliary information
- 2. Ratio estimator provides a precise estimate of the population mean if the regression is linear and passes through origin
- 3. Ratio estimator is biased
- A) 1 and 3 only B) 2 only C) 3 only D) 1, 2 and 3

53. If E(Y|X) = 1, then:

A)	$V(XY) \ge V(X)$	B)	$V(XY) \le V(X)$
C)	V(XY) = V(X)	D)	None of these

54. Which of the following is not a probability density function? A) f(x) = 1, 1 < x < 2 B) f(x) = x(2-x), 0 < x < 2C) $f(x) = 2, -\frac{1}{4} < x < \frac{1}{4}$ D) $f(x) = 1, -\frac{1}{2} < x < \frac{1}{2}$

- 55. The function defined by $F(x_1, x_2) = \begin{cases} 1, x_1 + 2x_2 \ge 1\\ 0, x_1 + 2x_2 < 1 \end{cases}$ represents:
 - A) The distribution function of discrete bivariate random variables X_1 and X_2
 - B) The distribution function of continuous bivariate random variables X_1 and X_2
 - C) The distribution function of bivariate random variables X_1 and X_2 of mixed type
 - D) Not a distribution function
- 56. Let X be an integer valued random variable with probability generating function (pgf) P(s). Then the pgf of 2X +1 is:
 - A) P(s) B) 2P(s) + 1 C) $sP(s^2)$ D) Does not exist

57. A random variable X has pdf f(x) = 1, 0 < x < 1. Then $P\left[\left|X - \frac{1}{2}\right| > \frac{1}{\sqrt{3}}\right] \leq \cdots$.

- A) 0.25 B) 0.05 C) 0.2 D) 0.01
- 58. If X is a random variable with $\beta_n = E|X|^n < \infty$, then for $2 \le k \le n$, which of the following is true?

A)
$$(\beta_{k-1})^{\frac{1}{k-1}} \le (\beta_k)^{\frac{1}{k}}$$
 B) $(\beta_{k-1})^{\frac{1}{k-1}} \ge (\beta_k)^{\frac{1}{k}}$

C)
$$(\beta_{k-1})^{\frac{1}{k-1}} = (\beta_k)^{\frac{1}{k}}$$
 D) $(\beta_{k-1})^{\frac{1}{k}} \le (\beta_{k-1})^{\frac{1}{k-1}}$

59. Let X_n be a random variable defined by P(X_n = n²) = 1/n and P(X_n = 0) = 1 - 1/n. Then which of the following is true?
A) X_n → 0 and E(X_n) → 0
B) X_n → 0 and E(X_n) → ∞
C) X_n → 1 and E(X_n) → 0
D) None of the above
60. Let {X_n} be a sequence of iid random variables with mean μ and finite variance σ². If S_n = X₁ + X₂ + ... + X_n, then P [lim_{n→∞} S_n/n = μ]=

- A) 0 B) $\frac{1}{2}$ C) 1 D) $\frac{\sigma^2}{\mu^2}$
- 61. Which of the following is not a characteristic function? A) e^{-t^4} B) $e^{-|t|^4}$ C) $(1 + t^4)^{-1}$ D) All of these
- 62. Let $X_1, X_2, ..., X_n$ be a random sample from $N(\mu, \sigma^2), \sigma^2$ is known. It is also known that $\mu \in (\theta_1, \theta_2), \theta_1 < \theta_2$. If \overline{X} is the sample mean and

$$T = \begin{cases} \theta_1, & \text{if } \bar{X} < \theta_1 \\ \bar{X}, & \text{if } \theta_1 \le \bar{X} \le \theta_2 \\ \theta_2, & \text{if } \bar{X} > \theta_2 \end{cases}$$

Then which of the following is true?

- A) *T* is unbiased estimator for μ with $MSE(\bar{X}) = MSE(T)$
- B) *T* is biased estimator for μ with $MSE(\bar{X}) < MSE(T)$
- C) T is unbiased estimator for μ with $MSE(\bar{X}) > MSE(T)$
- D) *T* is biased estimator for μ with $MSE(\bar{X}) > MSE(T)$
- 63. If $X_{(1)}$ and $X_{(n)}$ are the 1st and n^{th} order statistics of a random sample of size *n* from the rectangular distribution $U(a, \theta)$, where *a* is known. Then,
 - A) $X_{(n)}$ is sufficient for θ
 - B) $X_{(1)}$ and $X_{(n)}$ are jointly sufficient for θ
 - C) $\min(-X_{(1)}, X_{(n)})$ is sufficient for θ
 - D) $X_{(n)} X_{(1)}$ is sufficient for θ

64. Let X_1, X_2, \dots, X_n be a random sample from pdf $f(x) = \frac{1}{\beta} e^{-\left(\frac{x-\alpha}{\beta}\right)}, \quad \alpha < x < \infty, \quad -\infty < \alpha < \infty, \quad \beta > 0$ If $X_{(1)} = \min(X_1, X_2, \dots, X_n)$ and $\overline{X} = \frac{1}{n} \sum_{i=1}^n X_i$. Then the MLE of β is:

A)
$$X_{(1)}$$
 B) $\frac{1}{\bar{X} - X_{(1)}}$ C) $\bar{X} - X_{(1)}$ D) \bar{X}

- 65. Suppose X has density function $f(x) = \frac{1}{\sqrt{2\pi}}e^{-x^2/2}$, $x \in R$ under H_0 and $g(x) = \frac{1}{2}\exp\{-|x|\}$, $x \in R$ under H_1 . Then most powerful critical region for testing H_0 against H_1 based on a single observation is of the form:
 - A) |x| > k for a constant k
 - B) $|x| \le k$ for a constant k
 - C) $k_1 \le |x| \le k_2$ for some constant k_1 and k_2
 - D) $|x| \ge k_1$ or $|x| \le k_2$ for some constant k_1 and k_2
- 66. Which of the following family of pdf $\{f_{\theta}\}$ has monotone likelihood ratio property? 1. $f_{\theta}(x) = \frac{1}{2} \exp(-|x - \theta|), -\infty < x < \infty, \quad \theta \in \mathbb{R}$
 - 2. $f_{\theta}(x) = \frac{e^{-x-\theta}}{\left[1+e^{-x-\theta}\right]^2}, \quad -\infty < x < \infty, \quad \theta \in \mathbb{R}$ 3. $f_{\theta}(x) = \frac{1}{\theta}, \quad 0 < x < \theta, \quad \theta > 0$

A)	1 and 2 only	B)	1 and 3 only
C)	2 and 3 only	D)	1, 2 and 3

- 67. Let $X_1, X_2, ..., X_n$ be independent observations taken from the binomial distribution b(1, p). To test $H_0: p = \frac{1}{2}$ against $H_1: p = \frac{3}{4}$ sequential probability ratio test is applied. Let A and B, 0 < B < 1 < A be the boundary constants of the SPRT, then at the m^{th} stage, H_0 will be accepted if:
 - A) $\sum_{i=1}^{m} X_i \leq \log(2^m B)$ B) $\sum_{i=1}^{m} X_i \geq \log(2^m A)$

C)
$$\sum_{i=1}^{m} X_i \le \log \frac{B}{2^m}$$
 D) $\sum_{i=1}^{m} X_i \ge \log \frac{A}{2^m}$

- 68. Let $X_1, X_2, ..., X_m$ and $Y_1, Y_2, ..., Y_n$ be independent random samples from $N(\mu_1, \sigma_1^2)$ and $N(\mu_2, \sigma_2^2)$, μ_1, μ_2 are unknown. If $\lambda(\underline{x}, \underline{y})$ be the likelihood ratio used for testing $H_0: \sigma_1^2 = \sigma_2^2$ against $H_1: \sigma_1^2 \neq \sigma_2^2$. Then the asymptotic distribution of $-2\log\lambda(\underline{x}, \underline{y})$ is:
 - A) Chi-square with 1 degree of freedom
 - B) Chi-square with 2 degrees of freedom
 - C) Chi-square with m + n 1 degrees of freedom
 - D) Chi-square with m + n 2 degrees of freedom
- 69. The Fisher information in the sample of size 100 taken from the Poisson distribution with mean 4 is:
 - A) $\frac{1}{4}$ B) $\frac{1}{25}$ C) 5 D) 25

70. Let $X_1, X_2, ..., X_n$ be a random sample taken from $U(0, \theta)$. The uniformly most powerful test for testing $H_0: \theta = \theta_0$ against $H_1: \theta \neq \theta_0$ with level of significance α is given by:

$$\varphi(\underline{x}) = \begin{cases} 1, \ x_{(n)} > \theta_0 \text{ or } x_{(n)} < \theta_0 \alpha^{1/n} \\ 0, \text{ otherwise} \end{cases}$$

where $x_{(n)} = \max(X_1, X_2, \dots, X_n)$. Then the uniformly most accurate confidence interval for θ is:

- A) $[x_{(n)} \alpha^{1/n}, x_{(n)} + \alpha^{1/n}]$
- B) $\left[x_{(n)}\alpha^{1/n}, x_{(n)}\right]$
- C) $\left[x_{(n)}, x_{(n)}\alpha^{-1/n}\right]$
- D) Cannot be determined with the given information
- 71. Three brands of tea are rated for the taste on a scale of 1 to 10. Six persons are asked to rate each brand so that there is a total of 18 observations. The appropriate test to determine if three brand's taste equally good is:
 - A) One way analysis of variance
 - B) Friedman test
 - C) Kruskal-Wallis test
 - D) Wilcoxon rank-sum test
- 72. Let $X_1, X_2, ..., X_n$ be a random sample from a gamma distribution with PDF $f(x) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} e^{-\beta x} x^{\alpha-1}, x > 0, \alpha, \beta > 0$

If β is known, then which of the following is true?

A)	$\sum X_i$ is sufficient for α	B)	$\sum X_i^2$ is sufficient for α
C)	$\prod X_i$ is sufficient for α	D)	$(\sum X_i, \sum X_i^2)$ is sufficient for α

73. Under the assumptions required for the Wilcoxon's signed rank test, let T^+ be the sum of the ranks of positive X_i 's of a random sample $X_1, X_2, ..., X_n$. Then the asymptotic distribution of T^+ under the null hypothesis is:

A)
$$N\left(\frac{n(n+1)}{4}, \frac{n(n+1)(2n+1)}{12}\right)$$
 B) $N\left(\frac{n(n+1)}{2}, \frac{n(n+1)(2n+1)}{24}\right)$
C) $N\left(\frac{n(n+1)}{4}, \frac{n(n+1)(2n+1)}{24}\right)$ D) $N\left(\frac{n(n+1)}{6}, \frac{n(n+1)(2n+1)}{12}\right)$

- 74. Let $X_1, X_2, ..., X_n$, $n \ge 2$ be iid observations from $N(0, \sigma^2)$, where $\sigma^2 (> 0)$ is unknown. Then the UMVUE of σ^2 is:
 - A) $\frac{1}{n} \sum X_i^2$ B) $\frac{1}{n-1} \sum X_i^2$

C)
$$\frac{1}{n}\sum(X_i - \bar{X})^2$$
 D) $\frac{1}{n-1}\sum(X_i - \bar{X})^2$

75. A UMP test is:

-			
A)	Always exists	B)	Unbiased test
C)	Biased test	D)	None of these

- 76. Let $X \sim b(n, p)$ and p has the pdf $\pi(p) = 1$, $0 . Then the Baye's estimator of <math>p^2$ under the quadratic error loss function is:
 - A) $\frac{x}{n}$ B) $\frac{x(x+1)}{n(n+1)}$ C) $\frac{(x+1)(x+2)}{(n+2)(n+3)}$ D) $\frac{x^2}{n^2}$
- 77. Let $X_1, X_2, ..., X_n$ be a sequence of iid Bernoulli random variable with probability of success p. Also, let n be a positive integer valued random variable having Poisson distribution with mean θ . Then the mean and variance of $\sum_{i=1}^{n} X_i$ are:
 - A) Mean= θp , variance= θp
 - B) Mean= p, variance= $\theta^2 p^2$
 - C) Mean= $n\theta p$, variance= $n\theta p^2$
 - D) Mean= np, variance= $n\theta p(1-p)$
- 78. Let $X_1, X_2, ..., X_n$ be iid random variables, then which of the following statement is true:
 - A) $min(X_1, X_2, ..., X_n)$ has Weibull distribution if and only if the common distribution of X_i 's is Weibull.
 - B) If X_i 's are Weibull random variables, then $min(X_1, X_2, ..., X_n)$ is Weibull and the converse need not be true.
 - C) $min(X_1, X_2, ..., X_n)$ has Weibull distribution if and only if the common distribution of X_i 's is exponential.
 - D) All the above statements are wrong.

79. Let *X* has power series distribution with probability mass function

$$P(X = x) = \frac{a_x \theta^x}{A(\theta)}, x = 0, 1, 2, \dots; \theta > 0; a_j > 0; A(\theta) = \sum_{x=0}^{\infty} a_x \theta^x.$$

Then which of the following is true:

- A)variance(X) = mean(X)B)variance(X) = $\theta \frac{d}{d\theta}$ (mean)C)variance(X) = $\frac{d^2}{d\theta^2}$ (mean)D)variance(X) = $2\theta \frac{d^2}{d\theta^2}$ (mean)
- 80. For a set of *n* observations $Y_1, Y_2, ..., Y_n$, the maximum number of mutually orthogonal contrasts among them is:

A) n B) n-1 C) n-2 D) n-3